KIPAC Tea, 05/30/17 Oliver Friedrich (LMU gravitational lensing; Stella Seitz, Tamas Varga, Matthias Kluge++)

Precision matrix expansion – efficient use of numerical simulations in estimating errors on cosmological parameters

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- Data vectors & covariance matrices
- Multi-probe analyses (new orders of magnitude for covariance estimation)
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measurement of **ξ** as a function of **x**:

$$\hat{\xi}_1 \equiv \hat{\xi}(x_1) \ , \ \hat{\xi}_2 \equiv \hat{\xi}(x_2)$$

(unknown) expectation value of the measurement:

$$\xi_1 \equiv \langle \hat{\xi}_1 \rangle \,\,,\,\, \xi_2 \equiv \langle \hat{\xi}_2
angle$$

• covariance matrix of the measurement:

$$C_{ij} = \langle (\hat{\xi}_i - \xi_i) (\hat{\xi}_j - \xi_j) \rangle$$

$$\mathbf{C} = \begin{pmatrix} C_{11} & C_{12} \\ C_{21} & C_{22} \end{pmatrix}$$
$$= \begin{pmatrix} \sigma_1^2 & r\sigma_1\sigma_2 \\ r\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$$

(*r* = Pearson correlation)



- knowledge of the complete covariance needed to judge agreement between models and data
- Which model can be ruled out? (diagonal covariance)

 $\mathbf{C} = \begin{pmatrix} \sigma^2 & \mathbf{0} \\ \mathbf{0} & \sigma^2 \end{pmatrix}$



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- Which model can be ruled out? (diagonal covariance)

 $\mathbf{C} = \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix}$

 \rightarrow none



- knowledge of the complete covariance needed to judge agreement between models and data
- Which model can be ruled out? (strong positive correlation)

 $\mathbf{C} = \begin{pmatrix} \sigma^2 & 0.9\sigma^2 \\ 0.9\sigma^2 & \sigma^2 \end{pmatrix}$



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 \rightarrow model A



- knowledge of the complete covariance needed to judge agreement between models and data
- Which model can be ruled out? (strong negative correlation)

 $\mathbf{C} = \begin{pmatrix} \sigma^2 & -0.9\sigma^2 \\ -0.9\sigma^2 & \sigma^2 \end{pmatrix}$



- knowledge of the complete covariance needed to judge agreement between models and data
- Which model can be ruled out? (strong negative correlation)

$$\mathbf{C} = \begin{pmatrix} \sigma^2 & -0.9\sigma^2 \\ -0.9\sigma^2 & \sigma^2 \end{pmatrix}$$

 \rightarrow model B



Parameter inference:

assume that the data vector

 $\hat{\boldsymbol{\xi}} = \begin{pmatrix} \hat{\xi}_1 & \hat{\xi}_2 \end{pmatrix}$

has as multivariate Gaussian distribution:

$$p(\hat{\boldsymbol{\xi}}|\pi) \sim$$
 $\exp\left\{-\frac{1}{2}\left(\hat{\boldsymbol{\xi}} - \boldsymbol{\xi}[\pi]\right)^{T} \mathbf{C}^{-1}\left(\hat{\boldsymbol{\xi}} - \boldsymbol{\xi}[\pi]\right)
ight\}$

 neglect model parameters for which the data would be unlikely, i.e. :

$$p(\hat{\boldsymbol{\xi}}|\pi) < lpha$$

(Frequentist viewpoint)

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- Multi-probe analyses (new orders of magnitude for covariance estimation)
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Two point correlation functions:

- <u>galaxy clustering</u>: $w(\theta)$ (excess of galaxy pairs compared to a random distribution)
- <u>galaxy-galaxy lensing</u>: $\gamma_t(\theta)$ (tangential stretch of source images around lens galaxies)
- <u>cosmic shear</u>: $\xi_+(\theta)$, $\xi_-(\theta)$ (correlation of source galaxy shapes caused by lensing through matter inhomogeneities)





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Redshift tomography:

• split lens and source sample into several redshift bin

Ζ

look at cross-correlation

In our paper:

- <u>DES weak lensing only:</u> 450 data points
- <u>DES multi-probe:</u> 650 data points
- <u>LSST weak lensing only:</u> 2200 data points
 - → Want to get estimates of these covariances from N-body simulations!

Why is this difficult?

We need the inverse covariance: (the *precision matrix*)

$$\chi^{2}(\hat{\boldsymbol{\xi}}|\mathbf{C},\pi) = \left(\hat{\boldsymbol{\xi}} - \boldsymbol{\xi}[\pi]\right)^{\prime} \mathbf{C}^{-1} \left(\hat{\boldsymbol{\xi}} - \boldsymbol{\xi}[\pi]\right)$$

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constructing an unbiased estimator of the precision matrix:

$$\hat{\mathbf{C}}_{ij} = \frac{1}{N_s - 1} \sum_{k=1}^{N_s} (\hat{\xi}_i^k - \bar{\xi}_i) (\hat{\xi}_j^k - \bar{\xi}_j)$$
$$\hat{\boldsymbol{\Psi}} = \frac{N_s - N_d - 2}{N_s - 1} \hat{\mathbf{C}}^{-1}$$
$$\Rightarrow \langle \hat{\boldsymbol{\Psi}} \rangle = \mathbf{C}^{-1}$$

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estimation noisy of covariance and precision matrix:

$$\frac{\Delta \hat{C}_{ii}}{\hat{C}_{ii}} \approx \sqrt{\frac{2}{N_s - 1}}$$
$$\frac{\Delta \hat{\Psi}_{ii}}{\hat{\Psi}_{ii}} \approx \sqrt{\frac{2}{N_s - N_d - 2}}$$

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$$\frac{1}{1+x} = 1 - x + x^2 - \mathcal{O}(x^3)$$

1.) Avoid matrix inversion with a priori knowledge about the covariance

- true covariance: C
- covariance estimate: \hat{C}
- model covariance: M
- 'relative' deviation between model and true covariance:

 $\mathbf{X} = (\mathbf{C} - \mathbf{M}) \ \mathbf{M}^{-1}$

 $\mathbf{\hat{X}} = (\mathbf{\hat{C}} - \mathbf{M}) \ \mathbf{M}^{-1}$

Precision matrix expansion:

$$C = M + C - M$$
$$= (I + X) M$$

 $\Rightarrow \Psi \hspace{0.1in} = \hspace{0.1in} \mathsf{M}^{-1} \hspace{0.1in} [\mathsf{I} - \mathsf{X} + \mathsf{X}^2 - \mathcal{O}(\mathsf{X}^3)]$

• estimation of the first order term:

$$\mathbf{\hat{\Psi}}_{1st} = 2\mathbf{M}^{-1} - \mathbf{M}^{-1}\mathbf{\hat{C}} \ \mathbf{M}^{-1}$$

has noise

$$\sqrt{\frac{2}{N_s-1}}$$

 → much less noisy than standard estimator (but bias due to finite break of the series)

2.) Use simulations only for the covariance parts where you really need them

- constributions like shape-noise can accurately be modelled analytically!
- Split covariance as

C = A + B

where

A can be modelled accurately

and can be turned off in simulations

• let ${\sf B}_m$ be a model for ${\sf B}$ and ${\sf M}={\sf A}+{\sf B}_m$ the total model









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- let ${\sf B}_m$ be a model for ${\sf B}$ and

 $\mathbf{M}=\mathbf{A}+\mathbf{B}_m$ the total model

• Precision matrix expansion:

 $\Rightarrow \Psi = \mathsf{M}^{-1} \left[\mathsf{I} - \mathsf{X} + \mathsf{X}^2 - \mathcal{O}(\mathsf{X}^3) \right]$ with

$$\mathsf{X} = (\mathsf{B} - \mathsf{B}_m) \mathsf{M}^{-1}$$

• estimation of the first order term:

$$\hat{\boldsymbol{\Psi}}_{1\text{st}} = \boldsymbol{M}^{-1} - \boldsymbol{M}^{-1} \left(\hat{\boldsymbol{B}} - \boldsymbol{B}_m \right) \boldsymbol{M}^{-1}$$

 \rightarrow noisy part of the estimator becomes even smaller

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 estimation of the second order term: (with help of Letac and Massam 2004)

$$\hat{\Psi}_{2nd} = \mathbf{M}^{-1} + \mathbf{M}^{-1} \mathbf{B}_m \mathbf{M}^{-1} \mathbf{B}_m \mathbf{M}^{-1}$$
$$-\mathbf{M}^{-1} (\hat{\mathbf{B}} - \mathbf{B}_m) \mathbf{M}^{-1}$$
$$-\mathbf{M}^{-1} \hat{\mathbf{B}} \mathbf{M}^{-1} \mathbf{B}_m \mathbf{M}^{-1}$$
$$-\mathbf{M}^{-1} \mathbf{B}_m \mathbf{M}^{-1} \hat{\mathbf{B}} \mathbf{M}^{-1}$$
$$+\mathbf{M}^{-1} \frac{\nu^2 \hat{\mathbf{B}} \mathbf{M}^{-1} \hat{\mathbf{B}} - \nu \hat{\mathbf{B}} \operatorname{tr} (\mathbf{M}^{-1} \hat{\mathbf{B}})}{\nu^2 + \nu - 2} \mathbf{M}^{-1}$$

Results and open questions

- Even strong deviations of our fiducial covariance seems to yield a convergent series
- For DES cosmic shear the PME needs only 200 simulations to perform as the standard estimator with > 8000 sims
- slightly worse performance for DES multi-probe
- For LSST cosmic shear the PME needs only 2200 simulations to perform as the standard estimator with > 115.000 sims

- How do we know a priori, whether the series converges?
- Can we make use of other noise terms such as shot-noise?
- How do we account for the remaining additional scatter in best-fit parameters?

Show paper!

- Use halo model covariance by Krause & Eifler (2016) as
 C
- Deform it in different ways to produce a fake model covariance
 M

(through rescaling of some covariance parts & more complicated procedures)

Does the PME converge??



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How does it perform with finite number of sims??



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Conclusions

- PME robust towards strong deviations between model and N-body covariance
- for weak lensing only: excellent recovery of parameter constraints
- for galaxy clustering: still big improvement as opposed to standard way to estimate precision matrix





What if the PME does not converge??